

# Mathematics and numerics for data assimilation and state estimation – Lecture 13



Summer semester 2020

# Overview

- 1 Metropolis Hastings MCMC method
- 2 Smoothing in continuous state-space
  - Examples of dynamics
- 3 Well-posedness of smoothing
- 4 Smoothing for deterministic dynamics

## Summary of lecture 12

- Monte Carlo methods for sampling  $\pi$ :

$$\pi_{MC}^M[f] = \sum_{k=1}^M \frac{f(U_k)}{M}, \quad \text{where } U_k \stackrel{iid}{\sim} \pi$$

- Sampling the target (exactly or approximately)  $\pi$  indirectly through change of measure or an auxiliary/proposal distribution  $\hat{\pi}$ .
- Discrete-time continuous-space Markov chains
- Metropolis Hastings MCMC method.

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## Markov Chain Monte Carlo method (MCMC)

**Input:** target pdf  $\pi$ , a conditional proposal  $q(y|x)$  (i.e.,  $q(\cdot|x) \in \mathcal{M}$  for every  $x \in \mathbb{R}^d$ ).

**Output:** Markov chain  $X_0, X_1, \dots$  with objective that  $\pi_{MCMC}^M = \frac{1}{M} \sum_{k=1}^M \delta_{X_k}$  approximates measure associated to  $\pi$ .

### Metropolis-Hastings algorithm

Given  $X_n$ ,

1 generate proposal  $Y_n \sim q(\cdot|X_n)$

2 set

$$X_{n+1} = \begin{cases} Y_n & \text{with probability } \rho(X_n, Y_n) \\ X_n & \text{with probability } 1 - \rho(X_n, Y_n) \end{cases}$$

where the M-H acceptance probability is defined by

$$\rho(x, y) = \min \left( \frac{\pi(y) q(x|y)}{\pi(x) q(y|x)}, 1 \right)$$

# Assumptions and properties of Metropolis Hastings

## Assumptions

- must be able to sample from  $q(\cdot|x)$  for relevant  $x$
- $\pi$  must be known up to a constant (i.e., relevant for posterior densities with  $Z$  unknown),
- $q(\cdot|x)$  must be known up to a constant that is independent of  $x$ .

## Properties:

- When  $q(x|y) = q(y|x)$  the test ratio becomes

$$\frac{\pi(y) q(x|y)}{\pi(x) q(y|x)} = \frac{\pi(y)}{\pi(x)}.$$

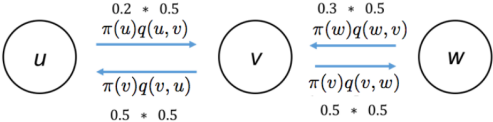
- If  $q(x|y) > q(y|x)$ , then (compared to not having a  $q$  ratio in the acceptance probability), the probability accepting transitions  $x \mapsto y$  increases. So transitions for which the reverse transition  $q(x|y)$  is more often proposed than the transition itself, increases likelihood.
- If  $q(x|y) < q(y|x)$ , then (compared to not having a  $q$  ratio in the acceptance probability), the probability accepting transitions  $x \mapsto y$  decreases.

# Effect of M-H acceptance

Top row: Markov chain with kernel density  $k(u, v) = q(v|u)$ .

Bottom row: M-H transforms kernel density to new kernel “density”

$$p(u, v) = \rho(u, v)q(v|u), \quad \text{with} \quad \rho(u, v) = \min \left( \frac{\pi(v)}{\pi(u)} \frac{q(u|v)}{q(v|u)}, 1 \right)$$



Let us modify the kernel:  
 $p(u, v) = a(u, v)q(u, v)$

↓

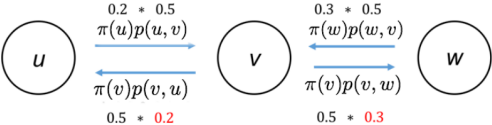
$$a(u, v) = \min \left( \frac{\pi(v)q(v, u)}{\pi(u)q(u, v)}, 1 \right)$$


Figure: From Data Assimilation and Inverse Problems, Sanz-Alonso et al.

M-H dynamics is associated to the transition kernel (ubung 5)

$$K(x, A) = \underbrace{\int_A \rho(u, v) q(y|x) dy}_{r(x, A)} + (1 - r(x, \mathbb{R}^d)) \delta_x(A)$$

Idea:

$$\begin{aligned} K(x, A) &= \mathbb{P}(X_1 \in A \mid X_0 = x) \\ &= \mathbb{P}(Y_0 \in A, X_1 = Y_0 \mid X_0 = x) + \mathbb{P}(x \in A, X_1 = x \mid X_0 = x) \end{aligned}$$



## M-H properties

If  $q(\cdot|x)$  dominates  $\pi$  for all  $x$ , then the M-H kernel satisfies detailed balance wrt  $\pi$ :

$$\int_A K(x, B)\pi(x)dx = \int_B K(x, A)\pi(x)dx \quad \forall A, B \in \mathcal{B}^d,$$

and  $\pi$  is an invariant pdf of the M-H Markov chain.

**Sketch of proof:** Assume that  $X_0 \sim \pi$ . Then

$$\begin{aligned}\mathbb{P}_1(A) &= \int_{\mathbb{R}^d} K(x, A)\mathbb{P}_{X_0}(dx) \\ &= \int_{\mathbb{R}^d} K(x, A)\pi(x)dx \\ &= \int_A K(x, \mathbb{R}^d)\pi(x)dx \\ &= \int_A \pi(x)dx = \mathbb{P}_0(A)\end{aligned}$$

## Remarks

**Challenges in real applications:** Choosing a proposal such that (1) one achieves convergence  $\pi^n \rightarrow \pi$ , (2) the convergence is fast in  $n$ , and (3) that acceptance of the proposal is frequent (for efficiency of MCMC).

See SST 6.4.2 for assumptions on prior and likelihood for  $\pi(\cdot|y)$  in combination with Gaussian proposal  $q(\cdot|x)$  which ensures convergence of the chain distribution.

If interested, “Monte Carlo Statistical Methods” by Robert and Casella is a good book on Monte Carlo and MCMC methods.

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## Dynamics and observation setting

**Continuous state-space dynamics:** A mapping  $\Psi \in C(\mathbb{R}^d, \mathbb{R}^d)$  is associated to the dynamics

$$\begin{aligned} V_{j+1} &= \Psi(V_j) + \xi_j, \quad j = 0, 1, \dots \\ V_0 &\sim N(m_0, C_0) \end{aligned} \tag{1}$$

where  $\{\xi_j\}$  is iid  $\xi \sim N(0, \Sigma)$ -distributed and  $V_0 \perp \{\xi_j\}$ .

**Observations:**

$$Y_j = h(V_j) + \eta_j, \quad j = 1, 2, \dots, \tag{2}$$

where  $h \in C(\mathbb{R}^d, \mathbb{R}^k)$  and  $\{\eta_j\}$  is iid with  $\eta_1 \sim N(0, \Gamma)$ .

**Independence assumptions:**

$$\{\eta_j\} \perp \{\xi_j\} \quad \text{and} \quad \{\eta_j\} \perp V_0.$$

**Objectives:** Study the smoothing pdf of  $V_{0:j} | Y_{1:j} = y_{1:j}$ .

## Examples of $\Psi$

In many applications,  $\Psi$  can be associated to a solution of a time-invariant ODE:

$$\begin{aligned}\dot{v} &= f(v), & t \geq 0 \\ v(0) &= v_0\end{aligned}\tag{3}$$

Viewing  $v_0$  as a **variable**, let us denote the solution of (3) at time  $s$  by  $\Psi(v_0; s)$ .

For a fixed interval  $\tau > 0$  and any  $V \in \mathbb{R}^d$ , we define

$$\Psi(V) := \Psi(V; \tau).$$

For later reference, let us also introduce

$$\Psi^{(j)}(V) := \underbrace{\Psi \circ \Psi \circ \dots \circ \Psi}_{j \text{ times}}(V) = \Psi(V; j\tau).$$

## Guiding examples

The scalar-valued ODE

$$\begin{aligned}\dot{v} &= \log(\lambda)v, & t \geq 0 \\ v(0) &= v_0\end{aligned}\tag{4}$$

and  $\tau = 1$  yields

$$\Psi(V) = e^{\log(\lambda)\tau} V = \lambda V.$$

The dynamics

$$V_{j+1} = \lambda V_j + \xi_j, \quad j = 0, 1, \dots$$

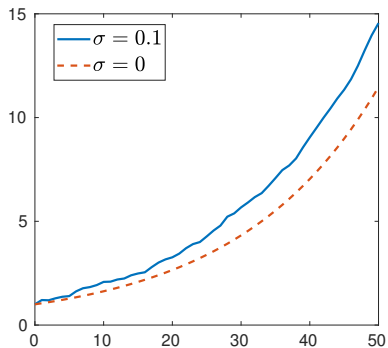
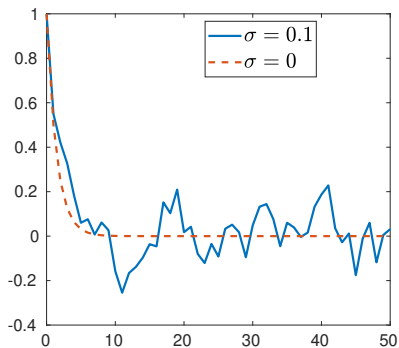
with  $\xi \sim N(0, \sigma^2)$  is fundamentally different when  $|\lambda| < 1$  and  $|\lambda| > 1$ .

Using that

$$\mathbb{E}[V_{j+1}] = \lambda \mathbb{E}[V_j], \quad \mathbb{E}[V_{j+1}^2] = \lambda^2 \mathbb{E}[V_j^2] + \sigma^2,$$

one can show that when  $|\lambda| < 1$ ,

$$\mathbb{P}_{V_n} \Rightarrow N\left(0, \frac{\sigma^2}{1 - \lambda^2}\right) \quad \text{as } n \rightarrow \infty.$$



**Figure:** Dynamics of  $V_{j+1} = \lambda V_j + \xi_j$  with  $\lambda = 0.5$  (left) and  $\lambda = 1.05$  (right).

## Nonlinear dynamics

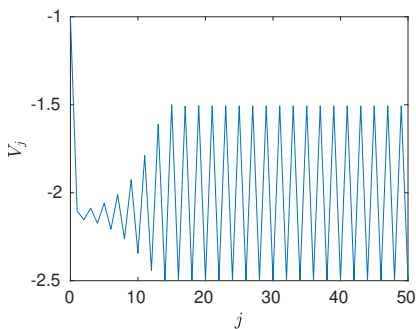
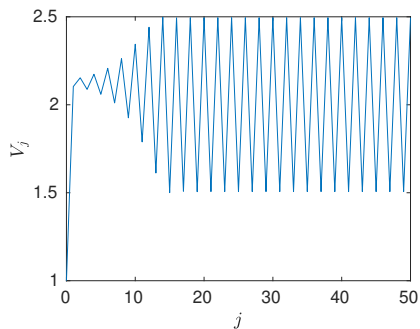
For

$$\Psi(v) = \alpha \sin(v)$$

the deterministic dynamics

$$V_{j+1} = \alpha \sin(V_j)$$

is sensitive to the initial condition.



**Figure:** Dynamics of with  $\alpha = 2.5$  and  $V_0 = 1$  (left) and  $V_0 = -1$  (right)



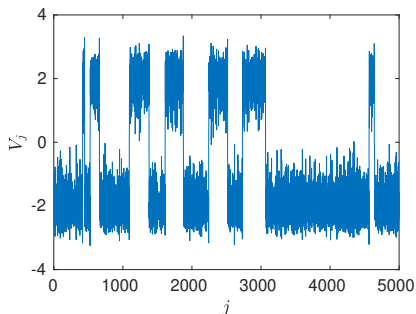
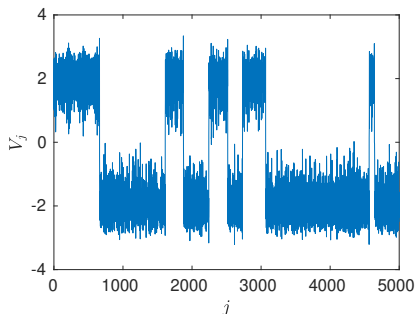
The stochastic dynamics

$$V_{j+1} = \alpha \sin(V_j) + \xi_j, \quad \xi \sim N(0, \sigma^2).$$

is not sensitive to the initial condition, if one views

$$\mathbb{P}_V(\cdot) := \lim_{J \rightarrow \infty} \frac{1}{J} \sum_{j=1}^J \delta_{V_j}(\cdot)$$

as the relevant feature (take as soft motivation, have not even shown that this measure exists).



**Figure:**  $\alpha = 2.5$ ,  $\sigma = 1/4$  and  $V_0 = 1$  (left) and  $V_0 = -1$  (right)

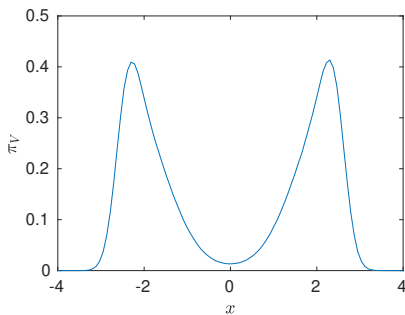
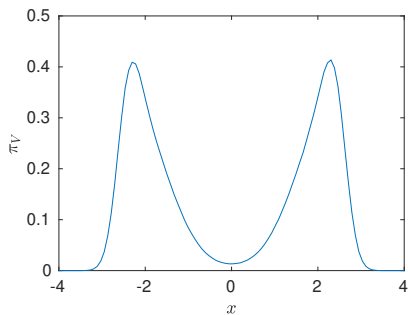
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as the relevant feature (take as soft motivation; we have not shown that this measure exists).



**Figure:**  $\pi_V = \text{PDF}(\mathbb{P}_V)$  for  $\alpha = 2.5$ ,  $\sigma = 1/4$ ,  $J = 10^7$ , and  $V_0 = 1$  (left) and  $V_0 = -1$  (right)

## Lorenz '63

Is the system of ODE

$$\left. \begin{aligned} \dot{v}_1 &= a(v_2 - v_1) \\ \dot{v}_2 &= -av_1 - v_2 - v_1v_3 \\ \dot{v}_3 &= v_1v_2 - bv_3 - b(r + a) \end{aligned} \right\} =: f(v), \quad t \geq 0,$$

where  $a, b, r > 0$  and  $v(0) \in \mathbb{R}^3$ .

For some  $\alpha, \beta > 0$ , depending on vector field, it can be shown that

$$f(v)^T v \leq \alpha - \beta|v|^2.$$

This ensures that [LSZ Example 1.22]

$$\limsup_{t \rightarrow \infty} |v(t)|^2 \leq \frac{\alpha}{\beta}$$

For any  $|v(0)| \leq \alpha/\beta$  there exists a unique solution, **see [übung 6](#)**, but  $v(t)$  is very sensitive to the initial condition!

Integration in Matlab with parameter values  $(a, b, r) = (10, 8/3, 28)$   
 $v(0) = (1, 1, 1)$  and  $\tilde{v}(0) = (1, 1, 1 + 10^{-5})$ :

```
options = odeset('RelTol',1e-12,'AbsTol',1e-10);
```

```
a = 10;
```

```
b = 8/3;
```

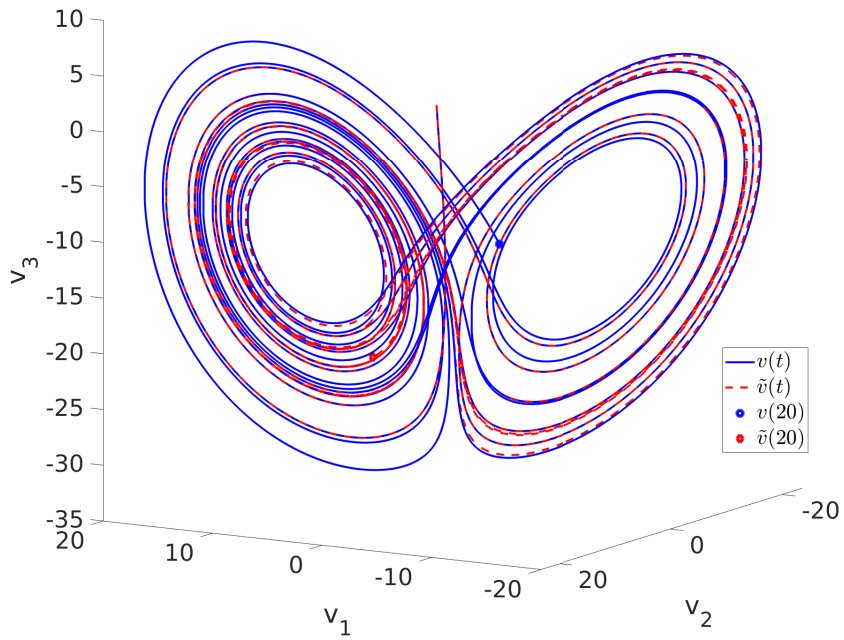
```
r = 28;
```

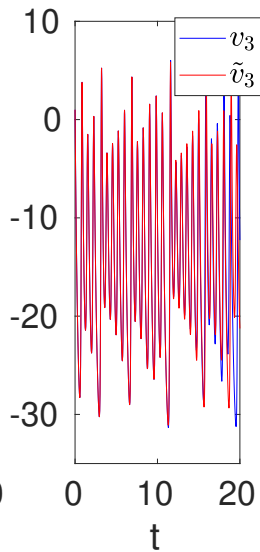
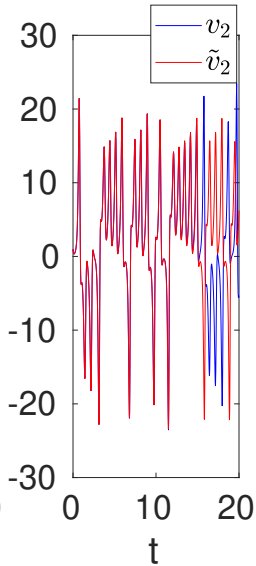
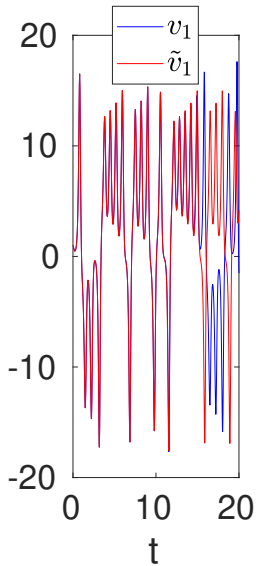
```
f = @(t,v) [a*(v(2)-v(1));  
            -a*v(1)-v(2)-v(1)*v(3);  
            v(1)*v(2)-b*v(3)-b*(r+a)];
```

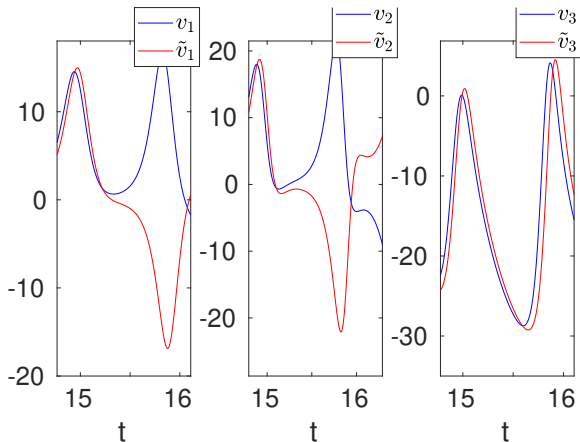
```
[t,v]=ode45(f,[0 20],[1 1 1], options);%RK 4/5 order ODE solve
```

```
[t2,vTilde] = ode45(f,[0 20],[1 1 1+1e-5], options);
```

Result:  $|v(0) - \tilde{v}(0)| = 10^{-5}$  and  $|v(20) - \tilde{v}(20)| \approx 15.7$







If  $v_1(s) = v_2(s) = 0$ , then  $(v_1, v_2) = (0, 0)$  for all later times: when  $v_3$  is sufficiently negative, it is an unstable stationary point on the  $(v_1, v_2)$ -subspace.

$$\dot{v}_1 = a(v_2 - v_1)$$

$$\dot{v}_2 = -av_1 - v_2 - v_1 v_3$$

$$\dot{v}_3 = v_1 v_2 - bv_3 - b(r + a)$$

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# Smoothing

Given dynamics

$$V_{j+1} = \Psi(V_j) + \xi_j, \quad \xi \sim N(0, \Sigma)$$
$$V_0 \sim N(m_0, C_0)$$

and observations

$$Y_j = h(V_j) + \eta_j, \quad \eta \sim N(0, \Gamma)$$

with  $h \in C(\mathbb{R}^d, \mathbb{R}^k)$  and  $V_0 \perp \{\eta_j\} \perp \{\xi_j\}$ .

**Objectives:** given  $y_{1:J} \in \mathbb{R}^{k \times J}$ ,

- derive the pdf for smoothing problem:

$$\pi_{V_{0:J}|Y_{1:J}}(v_{0:J}|y_{1:J}) =: \pi(v_{0:J}|y_{1:J})$$

- verify that the smoothing problem is stable wrt perturbations in  $y_{1:J} \in \mathbb{R}^{k \times J}$ . That is, show that

$$|y_{1:J} - \tilde{y}_{1:J}| = \mathcal{O}(\delta) \implies d_H(\pi(\cdot|y_{1:J}), \pi(\cdot|\tilde{y}_{1:J})) = \mathcal{O}(\delta)$$

## The smoothing pdf

By Bayes' rule and the Bayesian viewpoint

$$\pi(v_{0:J}|y_{1:J}) \propto \underbrace{\pi(y_{1:J}|v_{0:J})}_{\text{Likelihood}} \underbrace{\pi(v_{0:J})}_{\text{Prior}}$$

**Prior:** Note that  $\{V_j\}$  is a Markov chain, hence

$$\begin{aligned}\pi(v_{0:J}) &= \pi(v_J|v_{0:J-1})\pi(v_{0:J-1}) = \pi(v_J|v_{J-1})\pi(v_{0:J-1}) \\ &= \dots = \prod_{j=0}^{J-1} \pi(v_{j+1}|v_j)\pi_{v_0}(v_0).\end{aligned}$$

And

$$V_0 \sim N(m_0, C_0) \implies \pi_{v_0}(v_0) \propto \exp\left(-\frac{1}{2}|v_0 - m_0|_{C_0}^2\right),$$

and

$$\begin{aligned}V_{j+1}|(V_j = v_j) &= (\Psi(V_j) + \underbrace{\eta_j}_{\sim N(0, \Sigma)})|(V_j = v_j) \sim N(\Psi(v_j), \Sigma) \\ \implies \pi(v_{j+1}|v_j) &\propto \exp\left(-\frac{1}{2}|v_{j+1} - \Psi(v_j)|_{\Sigma}^2\right)\end{aligned}$$

**Prior:**

$$\pi(v_{0:J}) = \frac{1}{Z_P} \exp(-R(v_{0:J}))$$

where

$$R(v_{0:J}) := \frac{1}{2} |v_0 - m_0|_{C_0}^2 + \frac{1}{2} \sum_{j=0}^{J-1} |v_{j+1} - \Psi(v_j)|_{\Sigma}^2.$$

Next,

$$\pi(v_{0:J} | y_{1:J}) \propto \underbrace{\pi(y_{1:J} | v_{0:J})}_{\text{Likelihood}} \underbrace{\pi(v_{0:J})}_{\text{Prior}}$$

**Likelihood:** Since  $Y_j = h(V_j) + \eta_j$  and  $V_0 \perp \{\eta_j\} \perp \{\xi_j\}$ ,

$$\begin{aligned} Y_{1:J} | (V_{0:J} = v_{0:J}) &= (Y_1 | (V_1 = v_1), \dots, Y_J | (V_J = v_J)) \\ &= (h(v_1) + \eta_1, \dots, h(v_J) + \eta_J) \end{aligned}$$

with independent components and  $h(v_j) + \eta_j \sim N(h(v_j), \Gamma)$ .

Hence,

$$\pi(y_{1:J} | v_{0:J}) = \prod_{j=1}^J \pi(y_j | v_j) \propto \exp(-L(v_{1:J}; y_{1:J}))$$

with

$$L(v_{1:J}; y_{1:J}) := \frac{1}{2} \sum_{j=1}^J |h(v_j) - y_j|_{\Gamma}^2.$$

## Smoothing pdf

### Theorem 1

For the dynamics-observation sequence (1) and (2) with  $Y_{1:J} = y_{1:J}$ , we obtain

$$\begin{aligned}\pi(v_{0:J}|y_{1:J}) &= \frac{1}{Z} \exp(-L(v_{1:J}; y_{1:J}) - R(v_{0:J})) \\ &= \frac{1}{Z} \exp\left(-\frac{1}{2} \sum_{j=1}^J |h(v_j) - y_j|_{\Gamma}^2 \right. \\ &\quad \left. - \frac{1}{2} |v_0 - m_0|_{C_0}^2 - \frac{1}{2} \sum_{j=0}^{J-1} |v_{j+1} - \Psi(v_j)|_{\Sigma}^2\right)\end{aligned}$$

where  $v_{0:J} \in \mathbb{R}^{d \times (J+1)}$  and the normalizing constant  $Z$  depends on  $y_{1:J} \in \mathbb{R}^{k \times J}$

**Next question:** How stable is the pdf wrt perturbations in  $y_{1:J}$ ?

## Well-posedness of the smoothing pdf

### Theorem 2 (LSZ 2.15)

Fix  $J \in \mathbb{N}$ , a pair of observation sequences  $y_{1:J}, \tilde{y}_{1:J} \in \mathbb{R}^{k \times J}$ , and assume that the dynamics  $V_j$  satisfies

$$\mathbb{E} \left[ \sum_{j=0}^J (1 + |h(V_j)|^2) \right] < \infty.$$

Then there exists a constant  $c > 0$  that depends on  $y_{1:J}$  and  $\tilde{y}_{1:J}$  such that

$$d_H(\pi(\cdot | y_{1:J}), \pi(\cdot | \tilde{y}_{1:J})) \leq c \sqrt{\sum_{j=1}^J |y_j - \tilde{y}_j|^2}$$

## Proof ideas:

$$\pi(v_{0:J}|y_{1:J}) = \frac{1}{Z} \exp(-L(v_{1:J}; y_{1:J}) - R(v_{0:J}))$$

and

$$\pi(v_{0:J}|\tilde{y}_{1:J}) = \frac{1}{\tilde{Z}} \exp(-L(v_{1:J}; \tilde{y}_{1:J}) - R(v_{0:J}))$$

Results follows from showing that

$$Z, \tilde{Z} > K > 0 \quad \text{and} \quad |Z - \tilde{Z}| = \mathcal{O}(|y_{1:J} - \tilde{y}_{1:J}|)$$

and that

$$\begin{aligned} |L(v_{1:J}; y_{1:J}) - L(v_{1:J}; \tilde{y}_{1:J})| &= \frac{1}{2} \sum_{j=1}^J \left| |h(v_j) - y_j|_{\Gamma}^2 - |h(v_j) - \tilde{y}_j|_{\Gamma}^2 \right| \\ &= \mathcal{O}(|y_{1:J} - \tilde{y}_{1:J}|). \end{aligned}$$

Hint for bounding the loss-term difference: for  $u, v \in \mathbb{R}^k$ ,

$$|u|_{\Gamma}^2 - |v|_{\Gamma}^2 = \langle u + v, u - v \rangle_{\Gamma}$$

where  $\langle u, v \rangle_{\Gamma} := u^T \Gamma^{-1} v$ .

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## Smoothing problem – deterministic dynamics

Consider the simplified version of (1) where the dynamics is deterministic (but with random initial data):

$$\begin{aligned}V_{j+1} &= \Psi(V_j), & j = 0, 1, \dots, \\V_0 &\sim N(m_0, C_0)\end{aligned}$$

with observations  $j = 1, 2, \dots$

$$Y_j = h(V_j) + \eta_j, \quad \eta \sim N(0, \Gamma)$$

with  $h \in C(\mathbb{R}^d, \mathbb{R}^k)$  and  $V_0 \perp \{\eta_j\}$ .

Then, given  $Y_{1:J} = y_{1:J}$ , we now have that  $V_{0:J}$  **only is random in**  $V_0$ , since using that  $V_j = \Psi^{(j)}(V_0)$ ,

$$V_{0:J} = (V_0, \Psi(V_0), \Psi^{(2)}(V_0), \dots, \Psi^{(J)}(V_0)).$$

Consequently, we now seek to determine the pdf of  $V_0 | Y_{1:J} = y_{1:J}$ :

$$\pi(v_0 | y_{1:J}) \propto \underbrace{\pi(y_{1:J} | v_0)}_{\text{Likelihood}} \underbrace{\pi_{V_0}(v_0)}_{\text{Prior}}.$$

**Likelihood :** Since

$$Y_j = h(V_j) + \eta_j = h(\Psi^{(j)}(V_0)) + \eta_j$$

we obtain that

$$Y_{1:J}|(V_0 = v_0) = (h(\Psi^{(1)}(v_0)) + \eta_1, h(\Psi^{(2)}(v_0)) + \eta_2, \dots, h(\Psi^{(J)}(v_0)) + \eta_J).$$

This yields

$$\pi(y_{1:J}|v_0) = \prod_{j=1}^J \pi(y_j|v_0) \propto \exp \left( - \underbrace{\frac{1}{2} \sum_{j=1}^J \left| y_{j+1} - h(\Psi^{(j)}(v_0)) \right|_{\Gamma}^2}_{=L(v_0; y_{1:J})} \right)$$

and the posterior

$$\pi(v_0|y_{1:J}) \propto \exp \left( - L(v_0; y_{1:J}) - \underbrace{\frac{1}{2} |v_0 - m_0|_{C_0}^2}_{=R(v_0)} \right)$$

## Numerical study

For the dynamics

$$V_{j+1} = \lambda V_j$$

with  $V_0 \sim N(m_0, \sigma_0^2)$  and

$$Y_j = V_j + \eta_j, \quad \eta \sim N(0, \gamma^2)$$

it can be shown that

$$\pi(v_0 | y_{1:J}) \propto \exp \left( -\frac{1}{2\gamma^2} \sum_{j=1}^J |y_{j+1} - \lambda^j v_0|^2 - \underbrace{\frac{1}{2\sigma_0^2} |v_0 - m_0|^2}_{=R(v_0)} \right)$$

and completing squares in the exponent yields that

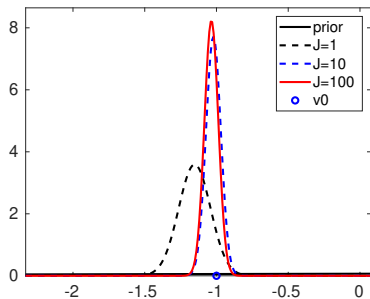
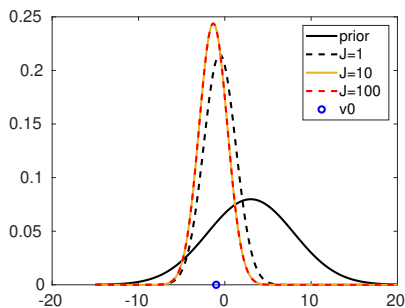
$$V_0 | (Y_{1:J} = y_{1:J}) \sim N(m, \sigma_{post}^2)$$

If  $|\lambda| < 1$ , then

$$\lim_{J \rightarrow \infty} \sigma_{post}^2 \stackrel{\text{a.s.}}{=} \frac{\gamma^2}{\lambda^2 / (1 - \lambda^2) + \gamma^2 / \sigma_0^2} \quad (\text{so uncertainty remains for large } J).$$

But cases when either  $\lambda^2 \approx 1$  and/or  $\gamma \approx 0$  reduce uncertainty (ubung 6).

Numerical test with  $\lambda = 1/2$ ,



**Figure:** Numerical tests with  $m_0 = 3$ ,  $\sigma_0 = 5$  from  $v_0 = -1$  and [**left**  $\lambda = 1/2$  and  $\gamma = 1$ ], [**right**  $\lambda = 0.9$  and  $\gamma = 0.1$ ].

See LSZ 2.8 for more illustrations of smoothing pdfs for  $V_0 | Y_{1:J}$ .

## Next time

We will talk about the filtering pdf  $\pi(v_j|y_{1:j})$  and Kalman filtering – i.e., filtering in the Gaussian-linear setting.